

Solution

Problem 1

- a. Since $E(B_t) = 0$ for all t , we clearly get $E(A_t) = 0$ as well. And since $E(B_s B_t) = s \wedge t$, we find $E(A_s A_t) = E(B_s B_t) - s E(B_1 B_t) - t E(B_s B_1) + s t E(B_1 B_1) = s \wedge t - s t - t s + s t = s \wedge t - s t$. Setting $s = t$ we get $E(A_t^2) = t - t^2$ as claimed.

We could further rewrite the answer as $E(A_s A_t) = (s \wedge t)(1 - s \vee t) = (s \wedge t)((1 - s) \wedge (1 - t))$, which makes the symmetry between A_t and A_{1-t} apparent.

- b. Clearly, with the new definition $E(A_t) = 0$. And since $t/(t-1)$ is a strictly increasing function of t , we get – assuming $s \leq t$ for simplicity

$$\begin{aligned} E(A_s A_t) &= (1-s)(1-t) E(B_{s/(1-s)} B_{t/(1-t)}) = (1-s)(1-t) \left(\frac{s}{1-s} \wedge \frac{t}{1-t} \right) \\ &= (1-s)(1-t) \frac{s}{1-s} = s(1-t) = s - s t = s \wedge t - s t \end{aligned}$$

Interchanging s and t yields the same result if $s \geq t$.

The distribution of variables within a Gaussian family is uniquely determined by expectations and covariances, so we conclude that (A_t) is a Brownian bridge.

For the converse, assume $(A_s)_{s \in [0,1]}$ is a Brownian bridge, and put $B_t = (1-s)^{-1} A_s$ where $t = s/(1-s)$. Solving for s we find $s = t/(1+t)$ and $(1-s)^{-1} = t+1$, so we end up with the definition

$$B_t = (1+t) A_{t/(1+t)} \quad t \geq 0.$$

Again, $E(B_t) = 0$ follows from $E(A_s) = 0$, and so – once more assuming $s \leq t$ for simplicity –

$$\begin{aligned} E(B_s B_t) &= (1+s)(1+t) E(A_{s/(1+s)} A_{t/(1+t)}) = (1+s)(1+t) \left(\frac{s}{1+s} \wedge \frac{t}{1+t} - \frac{s}{1+s} \frac{t}{1+t} \right) \\ &= (1+s)(1+t) \left(\frac{s}{1+s} - \frac{s}{1+s} \frac{t}{1+t} \right) = s = s \wedge t. \end{aligned}$$

It follows that (B_t) is a Brownian motion.

Problem 2

- a. A martingale with respect to some filtration (\mathcal{N}_t) is an (\mathcal{N}_t) -adapted process (M_t) with $E(|M_t|) < \infty$ satisfying $E(M_t | \mathcal{N}_s) = M_s$ when $s \leq t$.

The martingale representation theorem states that any L^2 martingale (M_t) (i.e., $E(M_t^2) < \infty$) with respect to the filtration (\mathcal{F}_t) given by n -dimensional Brownian motion (B_t) is given by an Itô integral:

$$M_t = M_0 + \int_0^t X_s dB_s \tag{1}$$

for a unique (\mathcal{F}_t) -adapted process (X_t) .

The converse follows from the basic theory of Itô integrals. Uniqueness follows directly from the Itô isometry.

- b. A stopping time τ with respect to a filtration (\mathcal{F}_t) is a random variable τ so that the event $\{\tau \leq t\}$ is measurable for all t .

In colloquial terms, it will be known at time t whether $\tau \leq t$ (“are we there yet?”) has occurred.

Now, for the stopped martingale $(M_{t \wedge \tau})$, we use the representation (1):

$$M_{t \wedge \tau} = M_0 + \int_0^{t \wedge \tau} X_s dB_s = M_0 + \int_0^t [s \leq \tau] X_s dB_s,$$

where $[s \leq \tau]$ equals 1 if $s \leq \tau$, and 0 otherwise. Since the process $([t \leq \tau] X_t)$ is clearly (\mathcal{F}_t) -adapted, the integral is a martingale, and we are done.

c. From the previous question, we get $E(M_{T \wedge t}) = E(E(M_{T \wedge t} | \mathcal{F}_0)) = E(M_{T \wedge 0}) = E(M_0)$. Whenever $\tau < \infty$, we clearly have $M_{T \wedge n} \rightarrow M_\tau$ as $n \rightarrow \infty$, and so $M_{T \wedge n} \rightarrow M_\tau$ a.s. Thanks to the boundedness assumption on $M_{T \wedge n}$, we can now employ the dominated convergence theorem to conclude that $E(M_\tau) = 0$.

d. Following the hint, we assume that $E(\tau) < \infty$.

Clearly also $B_{T \wedge t} \leq 1$, so $|M_{T \wedge t}^\alpha| \leq e^\alpha$, and we may conclude from point c that $E(M_\tau^\alpha) = E(M_0^\alpha) = 1$.

Since paths are continuous, $B_\tau = 1$, because the only way to exit $(-\infty, 1)$ is by passing through 1. Thus $M_\tau^\alpha = e^{\alpha - \alpha^2 \tau / 2}$, so we have in fact $E(e^{\alpha - \alpha^2 \tau / 2}) = 1$. Now we estimate:

$$1 = E(e^{\alpha - \alpha^2 \tau / 2}) \geq E(1 + \alpha - \alpha^2 \tau / 2) = 1 + \alpha - \frac{1}{2} \alpha^2 E(\tau),$$

which implies $\alpha E(\tau) \geq 2$ for any $\alpha > 0$. This contradicts the assumption $E(\tau) < \infty$.

Problem 3

a. Formally, the given equation can be written as

$$\frac{d^2 X_t}{dt^2} = -X_t + \frac{dB_t}{dt},$$

where the final term is a white noise term, which here plays the role of a random force acting on the mass.

Itô's formula yields

$$dU_t = X_t dX_t + Y_t dY_t + \frac{1}{2}((dX_t)^2 + (dY_t)^2) = X_t Y_t dt - Y_t X_t dt + Y_t dB_t + \frac{1}{2}(0 + dt) = Y_t dB_t + \frac{1}{2} dt.$$

Written on integral form:

$$U_t = U_0 + \frac{t}{2} + \int_0^t Y_s dB_s.$$

Since the expectation of an Itô integral is zero, we conclude $E(U_t) = E(U_0) + \frac{t}{2} = \frac{1}{2}(x^2 + y^2 + t)$.

b. We rewrite the equations as

$$dZ_t = HZ_t dt + \sigma dB_t, \quad Z_t = \begin{pmatrix} X_t \\ Y_t \end{pmatrix}, \quad H = \begin{pmatrix} 0 & 1 \\ -1 & 0 \end{pmatrix}, \quad \sigma = \begin{pmatrix} 0 \\ 1 \end{pmatrix}.$$

Using Itô's formula we get $d(e^{-tH} Z_t) = e^{-tH} (dZ_t - HZ_t dt) = e^{-tH} \sigma dB_t$, and so

$$e^{-tH} Z_t = Z_0 + \int_0^t e^{-sH} \sigma dB_s,$$

and therefore

$$Z_t = e^{tH} (x, y)^T + \int_0^t e^{(t-s)H} \sigma dB_s. \quad (2)$$

Since the integrand in the Itô integral is deterministic, we could certainly do a partial integration to turn the integral into an ordinary integral with a stochastic integrand instead, but this seems unnecessary for our purposes.

We focus on the matrix exponential (after verifying that $H^2 = -I$):

$$\begin{aligned} e^{tH} &= \sum_{n=0}^{\infty} \frac{(tH)^n}{n!} = \sum_{k=0}^{\infty} \frac{(tH)^{2k}}{(2k)!} + \sum_{k=0}^{\infty} \frac{(tH)^{2k+1}}{(2k+1)!} = \sum_{k=0}^{\infty} \frac{(-1)^k t^{2k}}{(2k)!} I + \sum_{k=0}^{\infty} \frac{(-1)^k t^{2k+1}}{(2k+1)!} H \\ &= (\cos t) I + (\sin t) H = \begin{pmatrix} \cos t & \sin t \\ -\sin t & \cos t \end{pmatrix}. \end{aligned}$$

On component form we can then write the solution as

$$\begin{aligned} X_t &= x \cos t + y \sin t + \int_0^t \sin(t-s) dB_s, \\ Y_t &= -x \sin t + y \cos t + \int_0^t \cos(t-s) dB_s. \end{aligned}$$

To compute $E(U_t)$, it may be better to stay with the vector formulation, noting that $U_t = \frac{1}{2}|Z_t|^2 = \frac{1}{2}Z_t^\top Z_t$. Plugging in (2), we note that when multiplying out the square, the cross term is another Itô integral, so it has expectation zero. Thus

$$E(U_t) = \frac{1}{2}|e^{tH}(x, y)^\top|^2 + \frac{1}{2}E\left(\left(\int_0^t e^{(t-s)H}\sigma dB_s\right)^2\right)$$

The first term is just $\frac{1}{2}|(x, y)^\top|^2 = \frac{1}{2}(x^2 + y^2)$, while the other term can be evaluated using the Itô isometry. The absolute value of the integrand turns out to be 1, so in the end we have

$$E(U_t) = \frac{1}{2}(x^2 + y^2 + t).$$

c. The infinitesimal generator is

$$A = y \frac{\partial}{\partial x} - x \frac{\partial}{\partial y} + \frac{1}{2} \frac{\partial^2}{\partial y^2}.$$

Since $u(t, x, y) = E^{(x,y)}(g(X_t, Y_t))$ where $g(x, y) = \frac{1}{2}(x^2 + u^2)$, this function should satisfy $\partial u / \partial t = Au$. With $u = \frac{1}{2}(x^2 + y^2 + t)$, we find $\partial u / \partial t = \frac{1}{2}$, while $Au = yx - xy + \frac{1}{2} = \frac{1}{2}$.